

## Today.

Johnson-Lindenstrass.

## Projections.

Project  $x$  into subspace spanned by  $v_1, v_2, \dots, v_k$ .

$$y_1 = x \cdot v_1, y_2 = x \cdot v_2, \dots, y_k = x \cdot v_k$$

Projection:  $(y_1, \dots, y_k)$ .

Have: Arbitrary vector, random  $k$ -dimensional subspace.

View As: Random vector, standard basis for  $k$  dimensions.

Orthogonal  $U$  - rotates  $v_1, \dots, v_k$  onto  $e_1, \dots, e_k$

$$y_i = \langle v_i | x \rangle = \langle Uv_i | Ux \rangle = \langle e_i | Ux \rangle = \langle e_i | z \rangle$$

Inverse of  $U$  maps  $e_i$  to random vector  $v_i$

$z = Ux$  is uniformly distributed on  $d$  sphere for unit  $x \in \mathbb{R}^d$ .

$y_i$  is  $i$ th coordinate of random vector  $z$ .

## Johnson-Lindenstrass

Points:  $x_1, \dots, x_n \in \mathbb{R}^d$ .

Random  $k = \frac{c \log n}{\epsilon^2}$  dimensional subspace.

Claim: with probability  $1 - \frac{1}{n^c}$ ,

$$(1 - \epsilon) \sqrt{\frac{k}{d}} |x_i - x_j| \leq |y_i - y_j| \leq (1 + \epsilon) \sqrt{\frac{k}{d}} |x_i - x_j|$$

"Projecting and scaling by  $\sqrt{\frac{d}{k}}$  preserves all pairwise distances w/in factor of  $1 \pm \epsilon$ ."

## Expected value of $y_i$ .

Random projection: first  $k$  coordinates of random unit vector,  $z_i$ .

$$E[\sum_{i \in [k]} z_i^2] = 1. \text{ Linearity of Expectation.}$$

By symmetry, each  $z_i$  is identically distributed.

$$E[\sum_{i \in [k]} z_i^2] = \frac{k}{d}. \text{ Linearity of Expectation.}$$

Expected length is (sort of)  $\sqrt{\frac{k}{d}}$ .

Johnson-Lindenstrass: close to expectation.

$k$  is large enough  $\rightarrow$

$$\approx (1 \pm \epsilon) \sqrt{\frac{k}{d}} \text{ with decent probability.}$$

## Random subspace.

Method 1:

Pick unit  $v_1$ ,

$v_2$  orthogonal to  $v_1$ ,

$\dots$

$v_k$  orthogonal to previous vectors...

Method 2:

Choose  $k$  vectors  $v_1, \dots, v_k$

Gram Schmidt orthonormalization of  $k \times d$  matrix where rows are  $v_i$ .  
remove projection onto previous subspace.

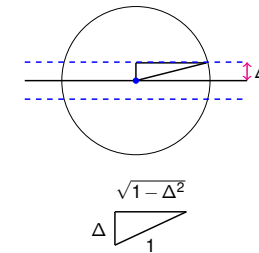
## Concentration Bounds.

$z$  is uniformly random unit vector.

Random point on the unit sphere.  $E[\sum_{i \in [k]} z_i^2] = \frac{k}{d}$ .

Claim:  $\Pr[|z_1| > \frac{t}{\sqrt{d}}] \leq e^{-t^2/2}$

Sphere view: surface "far" from equator defined by  $e_1$ .



$|z_1| \geq \Delta$  if  
 $z \geq \Delta$  from equator of sphere.  
Point on " $\Delta$ -spherical cap".

Area of caps

$\leq$  S.A. of sphere of radius  $\sqrt{1 - \Delta^2}$

$$\propto r^{d-1} = (1 - \Delta^2)^{(d-1)/2}$$

$$\propto \left(1 - \frac{t^2}{d}\right)^{(d-1)/2}$$

$$\propto \left(1 - \frac{t^2}{d}\right)^{\frac{d-1}{2}}$$

$$\approx e^{-\frac{t^2}{2}}$$

Constant of  $\propto$  is unit sphere area.  $\square$

$\Pr[\text{any } z_i^2 > (2 \log k) E[z_i^2]]$  is at most  $\approx 1/k$ .

## Many coordinates.

Argued  $\Pr[\text{any } z_i^2 > (2 \log k) E[z_i^2]]$  is at most  $\approx 1/k$ .

Total Length?  $z^2 = z_1^2 + z_2^2 + \dots + z_k^2$ .

Sum of ind. "gaussians" w/variance  $\sigma^2 = \sum_i \sigma_i^2 \leq k(1/d^2)$ .

$$\Pr\left[\left(z_1^2 + z_2^2 + \dots + z_k^2 - \frac{k}{d}\right) > t\right] \leq e^{-t^2/2\sigma^2} \leq e^{-t^2 d^2/2k}$$

Substituting  $t = \varepsilon \frac{k}{d}$ ,  $k = \frac{c \log n}{\varepsilon^2}$  and

$$\Pr\left[\left|z_1^2 + z_2^2 + \dots + z_k^2 - \frac{k}{d}\right| > \varepsilon \frac{k}{d}\right] \leq e^{-\varepsilon^2 k/2} = e^{-c \log n/2} = \frac{1}{n^{\varepsilon^2/2}}$$

Squared length within  $1 + \varepsilon$  of expectation, so length is too.

Roughly:  $\sqrt{1 + \varepsilon} \approx 1 + \varepsilon/2$ .

**Johnson-Lindenstraus:** For  $n$  points,  $x_1, \dots, x_n$ , all distances preserved to within  $1 \pm \varepsilon$  under  $\sqrt{\frac{k}{d}}$ -scaled projection above.

View one pair  $x_i - x_j$  as vector. Scale to unit.

Projection fails to preserve  $|x_i - x_j|$  with probability  $\leq \frac{1}{n^c}$

$\leq n^2$  pairs plus union bound

$\rightarrow$  prob any pair fails to be preserved with  $\leq \frac{1}{n^{c-2}}$ .

## Implementing Johnson-Lindenstraus

Random vectors have many bits

Use random bit vectors:  $\{-1, +1\}^d$  instead.

Almost orthogonal.

Project  $z$ .

Coordinate for bit vector  $b$ .

$$C_i = \frac{1}{\sqrt{d}} \sum_j b_j z_j$$

$$E[C_i^2] = E\left[\frac{1}{d} \sum_{i,j} b_i b_j z_i z_j\right] = \frac{1}{d} \sum_{i,j} E[b_i b_j] z_i z_j = \frac{1}{d} \sum_i z_i^2 = \frac{1}{d}$$

$$E[\sum_i C_i^2] = \frac{k}{d}$$

## Locality Preserving Hashing

Find nearby points in high dimensional space.

Points could be images!

Hash function  $h(\cdot)$  s.t.  $h(x_i) = h(x_j)$  if  $d(x_i, x_j) \leq \delta$  with good probability.

Low dimensions:

grid cells with sidelength  $\ell$

give additive  $\pm \sqrt{d} \ell$ -approximation.

Not quite a solution. Why?

Close to grid boundary.

Find close points to  $x$ :

Check grid cell and neighboring grid cells.

Project high dimensional points into low dimensions.

Use grid hash function for each coordinate.

## Binary Johnson-Lindenstraus

Project onto  $[-1, +1]$  vectors.

$$E[C] = E[\sum_i C_i^2] = \frac{k}{d}$$

Concentration?

$$\Pr\left[\left|C - \frac{k}{d}\right| \geq \varepsilon \frac{k}{d}\right] \leq e^{-\varepsilon^2 k}$$

Choose  $k = \frac{c \log n}{\varepsilon^2}$ .

$\rightarrow$  failure probability  $\leq 1/n^c$ .

## Hamming Distance.

Points are bit vectors with  $d$  bits.

Hamming distance in  $d$ -dimensions.

Is  $d(x, y) = \Delta$ ?

Idea: sample  $d/\Delta$  bits.

With constant probability you "miss" all differences.

$$\Pr[\text{miss}] = (1 - \Delta/d)^{d/\Delta} \approx 1/e$$

With constant probability have (odd) number of differences.

XOR the bits in sample to build a "coordinate":  $c(x)$ .

Property:

$$\Pr[c(x) \neq c(y) | d(x, y) \leq \Delta/2] < \Pr[c(x) \neq c(y) | d(x, y) \geq \Delta]$$

$O(\log n)$  coordinates. Law of large numbers.

Note: is not linear in difference.

Thus, data structure for each value of  $\Delta$ .

Property:

$$\Pr[c(x) \neq c(y) | d(x, y) \leq \Delta] < \Pr[c(x) \neq c(y) | d(x, y) \geq (1 + \varepsilon)\Delta]$$

## Analysis Idea.

$$\Pr\left[\left|C - \frac{k}{d}\right| \geq \varepsilon \frac{k}{d}\right] \leq e^{-\varepsilon^2 k}$$

Variance of  $C^2$ ? Recall  $C_i = \frac{1}{\sqrt{d}} \sum_j b_j z_j$

$$\text{Var}(C) \leq \left(\frac{k}{d^2}\right) (\sum_j z_j^4 + 4 \sum_{i,j} z_i^2 z_j^2) \leq \left(\frac{k}{d^2}\right) 2(\sum_i z_i^2)^2 \leq \frac{2k}{d^2}$$

Roughly normal (gaussian):

Density  $\propto e^{-t^2/2}$  for  $t$  std deviations away.

So, assuming normality

$$\sigma = \frac{\sqrt{2k}}{d}, \quad t = \frac{\varepsilon \frac{k}{d}}{\frac{\sqrt{2k}}{d}} = \varepsilon \sqrt{k}/\sqrt{2}$$

Probability of failure roughly  $\leq e^{-t^2/2}$

$\rightarrow e^{-\varepsilon^2 k/4}$

"Roughly normal." Chernoff, Berry-Esseen, Central Limit Theorems.

## Summary

Johnson-Lindenstrass.

$O(\log n)$  dimensions give good approximation of distances.

Projection onto  $k$  dense vectors.

Necessary as input vector may  $(1, 0, 0, 0, \dots)$ .

Compared with  $(1/\sqrt{n}, 1/\sqrt{n}, \dots, 1/\sqrt{n})$ .

Sampling a few coordinates of the latter works!!

Fast Johnson-Lindenstrass:

FFT transforms sparse vector into dense vector.

Also, transforms most dense vectors into dense vectors.

Premultiply by  $+/- 1$  diagonal matrix.

FFT:  $O(d \log d)$  to compute.

Can reduce  $O(d \log n / \epsilon^2)$  to  $O(d \log d + (\log^2 n) / \epsilon^3)$ .