CS 287, Fall 2011 Problem Set #2
Multivariate Gaussians, Kalman Filtering, Maximum Likelihood, EM

Deliverable: (1) Reasonable number of pages write-up in pdf format (2) zip file with your source code and clear “main” file for each programming related question. Due date/time: Thursday November 10th, 23:59pm. Email to pabbeel@berkeley.edu.

Please refer to the class webpage for the homework policy.
Various starter files are provided on the course website.
When making your write-up, make sure to answer all questions, and include and discuss plots (and if helpful, snippets of code) which are helpful in demonstrating that your system works and in answering the questions.

1. Maximum Likelihood
The Poisson distribution is a discrete probability distribution that expresses the probability of a number of events occurring in a fixed period of time if these events occur with a known average rate and independently of the time since the last event. Assume we obtain \( m \) i.i.d. samples \( x^{(0)}, ..., x^{(m)} \) distributed according to the Poisson distribution \( P(x = k) = \frac{\lambda^k e^{-\lambda}}{k!} \) for \( k = 1, 2, 3, ... \). What is the maximum likelihood estimate of \( \lambda \) as a function of \( x^{(0)}, ..., x^{(m)} \)?

2. Linearity of Expectation, Positive Semi-definiteness
A matrix \( A \in \mathbb{R}^{n \times n} \) is positive semi-definite (often denoted by \( A \succeq 0 \)) if and only if:

\[
A_{ij} = A_{ji} \\
\forall z \in \mathbb{R}^n : z^\top Az \geq 0
\]

Prove that covariance matrices, i.e., matrices of the form \( \Sigma = E[(X - EX)(X - EX)^\top] \) are positive semi-definite.

3. Linear Regression.
(a) Recursive Least Squares and Kalman Filtering. In recursive least squares one tries to solve the following problem: Let

\[
X^{(i)} = \begin{bmatrix} x^{(1)} \\ x^{(2)} \\ \vdots \\ x^{(i)} \end{bmatrix} \in \mathbb{R}^{i \times n}, Y^{(i)} = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(i)} \end{bmatrix} \in \mathbb{R}^{i \times 1}.
\]

One is asked to solve a sequence of least squares problems of the form \( \theta^{(i)} = \arg \min_{\theta} \|X^{(i)}\theta - Y^{(i)}\|_2^2 \), for \( i = 1, 2, ... \). This problem comes up when we incrementally get training examples
\((x^{(i)}, y^{(i)})\) and are asked to perform a least squares estimate as the data comes in. A standard application is one where \(\theta\) corresponds to the parameters of a communication channel (modeled as a linear system).

It turns out it is possible to solve this sequence of problems in an incremental fashion. One way to derive the algorithm is to perform a bunch of matrix algebra. A, certainly from our course background, simpler way to derive the incremental updates is to show how this corresponds to a Kalman filtering problem. Provide this derivation and then implement the corresponding Kalman filter and run it over the data provided in p3_a_starter.m. Plot the entries of \(\theta^{(i)}\) as a function of \(i\).

Hint: let \(x^{(i)} \in \mathbb{R}^n\). As there is no prior over \(\theta\) (or equivalently, a degenerate prior that is equal for all \(\theta \in \mathbb{R}^n\)) the first time you’ll be able to get a meaningful posterior density for \(\theta\) is when you have seen \(n\) datapoints. Your solution will start by first finding the posterior \(P(\theta | (x^{(1)}, y^{(1)}), (x^{(2)}, y^{(2)}), \ldots, (x^{(n)}, y^{(n)})) \propto \prod_{i=1}^n P(y^{(i)} | x^{(i)}, \theta)\). From then onwards an incremental update (which you can derive from your knowledge about Kalman filtering) will be performed to find \(P(\theta | (x^{(1)}, y^{(1)}), \ldots, (x^{(t+1)}, y^{(t+1)}))\) from \(P(\theta | (x^{(1)}, y^{(1)}), \ldots, (x^{(t)}, y^{(t)}))\).

(b) Bayesian Linear Regression. [This question does not attempt to build on top of the previous one and no incremental solutions should be used for this one.]

In lecture we saw how least squares can be seen as a maximum likelihood estimation of \(\theta\) for data assumed to come from \(y^{(i)} \sim \mathcal{N}(x^{(i)} \cdot \theta, \sigma^2)\). In this question we consider the Bayesian setting. In the Bayesian setting one treats \(\theta\) as a random variable. We assume a prior distribution: \(\theta \sim \mathcal{N}(0, \Sigma_0)\). When presented with \((x^{(i)}, y^{(i)})\) we assume that \(y^{(i)} \sim \mathcal{N}(x^{(i)} \cdot \theta, \sigma^2)\). We’ll assume \(\sigma\) is non-random and known.

Find an expression for the posterior over \(\theta\) after having seen \(k\) samples, \(P(\theta | (x^{(1)}, y^{(1)}), (x^{(2)}, y^{(2)}), \ldots, (x^{(k)}, y^{(k)}))\). In particular, show that the resulting distribution is a normal distribution and provide a formula for its mean and covariance.

Consider the posterior covariance and discuss what samples \(x^{(i)}\) tend to be most informative? (i.e., tend to result in smaller posterior covariance)

Given a new data point \(x\) and the posterior for \(\theta\), what is the distribution for \(x \cdot \theta\)?

Implement Bayesian linear regression following the instructions in p3_b_starter.m.

4. Kalman Filtering, Smoothing, EM

(a) Implementation of KF, Smoothing, EM. In this question you will implement a Kalman Filter, a Kalman Smoother, and the EM algorithm to estimate the covariance matrices. Look at p4_a_starter.m for more detailed instructions.

(b) Application to Species Population Size Estimation from Observations of Total Population Size. Consider three species \(U, V, W\) that grow independently of each other, exponentially with growth rates: \(U\) grows 2\% per hour, \(V\) grows 6\% per hour, and \(C\) grows 11\% per hour. The goal is to estimate the initial size of each population based on the measurements of total population.
Let \( x_U(t) \) denote the population size of species \( U \) after \( t \) hours, for \( t = 0, 1, \ldots \), and similarly for \( x_V(t) \) and \( x_W(t) \), so that
\[
x_U(t + 1) = 1.02x_U(t), \quad x_V(t + 1) = 1.06x_V(t), \quad x_W(t + 1) = 1.11x_W(t).
\]
The total population measurements are \( y(t) = x_U(t) + x_V(t) + x_W(t) + v(t) \), where \( v(t) \) are IID, \( \mathcal{N}(0, 0.36) \). (Thus the total population is measured with a standard deviation of 0.6).

The prior information is that \( x_U(0), x_V(0), x_W(0) \) (which we want to estimate) are IID \( \mathcal{N}(6, 2) \). (Obviously the Gaussian model is not completely accurate since it allows the initial populations to be negative with some small probability, but we’ll ignore that.)

How long will it be (in hours) before we can estimate \( x_U(0) \) with a variance less than 0.01? How long for \( x_V(0) \)? How long for \( x_W(0) \)?

(c) Correlated Noise. In many practical situations the noise is not independent. Consider the following stochastic system, for which the noise is not independent:
\[
x_0 \sim \mathcal{N}(\mu_0, \Sigma_0)
\]
\[
x_{t+1} = Ax_t + w_t
\]
\[
w_t = 0.3w_{t-1} + 0.2w_{t-2} + p_{t-1}
\]
\[
p_t \sim \mathcal{N}(0, \Sigma_{pp})
\]
\[
y_t = Cx_t + v_t
\]
\[
v_t = 0.8v_{t-1} + q_{t-1}
\]
\[
q_t \sim \mathcal{N}(0, \Sigma_{qq})
\]
\[
p_{t-1} = q_{t-1} = v_{-1} = w_{-1} = w_{-2} = 0
\]

Describe how, by choosing the appropriate state representation, the above setup can be molded into a standard Kalman filtering setup. In particular, describe the state, the dynamics model, and the measurement model such that the problem is transformed into the standard Kalman filtering setup with uncorrelated noise.

(d) (Optional / Extra Credit) EM Equations for \( A, B, C, d \). Derive the EM update equations for \( A, B, D, d \) for the usual linear Gaussian system, which is of the form:
\[
x_{t+1} = Ax_t + Bu_t + w_t \sim \mathcal{N}(0, \Sigma_w)
\]
\[
y_t = Cx_t + d + v_t \sim \mathcal{N}(0, \Sigma_v)
\]
where all \( w_t \) and \( v_t \) are independent. Show your work. Generate some data from a linear Gaussian system and report on the ability to learn \( A, B, C, d \) using EM.

5. Sensor Selection
We consider the following linear system:
\[
x_{t+1} = Ax_t + w_t
\]
\[
z_t = C_t x_t + v_t
\]
where \( A \in \mathbb{R}^{n \times n} \) is constant, but \( C_t \) can vary with time. The noise contributions are independent, and
\[
x_0 \sim \mathcal{N}(0, \Sigma_0), \quad w_t \sim \mathcal{N}(0, \Sigma_w) \quad v_t \sim \mathcal{N}(0, \Sigma_v).
\]
Here is the twist: the measurement matrix $C_t$ at each time comes from the set $\mathcal{S} = \{S_1, \ldots, S_K\}$. In other words, at each time $t$, we have $C_t = S_i_t$. The sequence $i_0, i_1, i_2, \ldots$ specifies which of the $K$ possible measurements is taken at time $t$. For example the sequence $2, 2, 2, \ldots$ means that $C_t = S_2$ for all $t$. The sequence $1, 2, \ldots, K, 1, 2, \ldots, K, \ldots$ is called round-robin: we cycle through the possible measurements, in order, over and over again.

Here is the interesting part: you get to choose the measurement sequence $i_0, i_1, i_2, \ldots$.

You will work with the following specific system:

$$
A = \begin{bmatrix}
1.0007 & -0.0010 & 0.0160 \\
0.0112 & 0.9944 & 0.0077 \\
-0.0003 & -0.0062 & 1.0009
\end{bmatrix}, \quad 
\Sigma_w = \text{diag}([0.1 \ 0.1 \ 1.0]), \quad \Sigma_v = 0.1^2, \quad \Sigma_0 = I
$$

and $K = 3$ with

$$
S_1 = [1.0 \ 1.0 \ 0], \quad S_2 = [0.2 \ 0.2 \ 0], \quad S_3 = [0 \ 0 \ 0.1].
$$

(a) **Using One Sensor.** Plot $\text{trace}(\Sigma_{t|0:t})$ versus $t$ for the three special cases when $C_t = S_1$ for all $t$, $C_t = S_2$ for all $t$, and $C_t = S_3$ for all $t$.

(b) **Round-robin.** Plot $\text{trace}(\Sigma_{t|0:t})$ versus $t$ using the round-robin sensor sequence $1, 2, 3, 1, 2, 3, \ldots$.

(c) **Greedy Sensor Selection.** Plot $\text{trace}(\Sigma_{t|0:t})$ versus $t$ using greedy sensor selection. In greedy sensor selection at time $t$ the choice of $i_0, i_1, \ldots, i_{t-1}$ has already been made and it has determined $\Sigma_{t-1|0:t-1}$. Then $\Sigma_{t|0:t}$ depends on $i_t$ only, i.e., which of $S_1, \ldots, S_K$ is chosen as $C_t$. Among these $K$ choices you pick the one that minimizes $\text{trace}(\Sigma_{t|0:t})$.

(d) **(Optional / Extra Credit) h-Lookahead Sensor Selection.** In $h$-lookahead sensor selection, at time $t$ the choice of $i_0, i_1, \ldots, i_{t-1}$ has already been made and it has determined $\Sigma_{t-1|0:t-1}$. Then $\Sigma_{t+h|0:t-1+h}$ depends on $i_t, i_{t+1}, \ldots, i_{t-1+h}$. A search over all possible sensor selection sequences for $t, t+1, \ldots, t-1+h$ is run, and let’s say $i^*_t, i^*_{t+1}, \ldots, i^*_t$ is the one that minimizes $\text{trace}(\Sigma_{t+h|0:t-1+h})$, then we choose $i_t = i^*_t$, and then we increment $t$ by one and repeat the process. Note that greedy sensor selection corresponds to $h=1$. Plot $\text{trace}(\Sigma_{t|0:t})$ for $h = 1, 2, \ldots$ (however far your algorithm is able to look ahead in a reasonable amount of time).

In all parts show the plots over the interval $t = 0, \ldots, 50$ and report the steady-state ($t \to \infty$) values (if such a limit exists).

Note none of these require knowledge of the measurements $z_0, z_1, \ldots$.

6. **(Optional / Extra Credit) EKF and UKF for Tracking, Localization, and SLAM**

In this question you build an EKF and a UKF for a car in the following contexts: (a) GPS and heading measurements are available. (b) GPS measurements are available. (c) Landmarks for which the locations are known are observed. (d) Landmarks for which only the ID is known are observed — this is E/UKF SLAM. (e) Landmarks are observed but neither their location nor their ID are available.

See p6_starter.m for detailed instructions. Make sure to discuss plots you generate a little bit—as you would do in a paper for experiments you ran and included in the paper.