

An important operation involving multivariate Gaussian distribution is the block matrix inversion. First, define the Schur complement of the matrix M with respect to H as

$$M/H = E - FH^{-1}G \quad (1)$$

where

$$M = \begin{pmatrix} E & F \\ G & H \end{pmatrix} \quad (2)$$

It can be verified that

$$\begin{pmatrix} I & -FH^{-1} \\ 0 & I \end{pmatrix} \begin{pmatrix} E & F \\ G & H \end{pmatrix} \begin{pmatrix} I & 0 \\ -H^{-1}G & I \end{pmatrix} = \begin{pmatrix} E - FH^{-1}G & 0 \\ 0 & H \end{pmatrix} \quad (3)$$

Then,

$$M^{-1} = \begin{pmatrix} I & 0 \\ -H^{-1}G & I \end{pmatrix} \begin{pmatrix} (M/H)^{-1} & 0 \\ G & H^{-1} \end{pmatrix} \begin{pmatrix} I & -FH^{-1} \\ 0 & I \end{pmatrix} \quad (4)$$

$$= \begin{pmatrix} (M/H)^{-1} & -(M/H)^{-1}FH^{-1} \\ -HG(M/H)^{-1} & H^{-1} + H^{-1}G(M/H)^{-1}FH^{-1} \end{pmatrix} \quad (5)$$

From the above expression, it can be verified that

$$|M| = |M/H||H| \quad (6)$$

and also,

$$(M/E)^{-1} = (H - GE^{-1}F)^{-1} = H^{-1} + H^{-1}G(E - FH^{-1}G)^{-1}FH^{-1} \quad (7)$$

Matrix inversion lemma:

Let

$$X = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \quad \mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} \quad \Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \quad (8)$$

Then,

$$\exp\left\{-\frac{1}{2}(X - \mu)^T \Sigma^{-1}(X - \mu)\right\} \quad (9)$$

$$= \exp\left\{-\frac{1}{2} \begin{pmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{pmatrix}^T \begin{pmatrix} I & 0 \\ -\Sigma_{22}^{-1}\Sigma_{21} & I \end{pmatrix} \begin{pmatrix} (\Sigma/\Sigma_{22})^{-1} & 0 \\ 0 & \Sigma_{22}^{-1} \end{pmatrix} \begin{pmatrix} I & -\Sigma_{21}\Sigma_{22}^{-1} \\ 0 & I \end{pmatrix} \begin{pmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{pmatrix}\right\} \quad (10)$$

$$= \exp\left\{-\frac{1}{2}(x_1 - \mu_1 - \Sigma_{12}\Sigma_{22}^{-1}(x_2 - \mu_2))^T (\Sigma/\Sigma_{22})^{-1}(x_1 - \mu_1 - \Sigma_{12}\Sigma_{22}^{-1}(x_2 - \mu_2))\right\} \\ \cdot \exp\left\{-\frac{1}{2}(x_2 - \mu_2)^T \Sigma_{22}^{-1}(x_2 - \mu_2)\right\} \quad (11)$$

The normalization can be split into two factors:

$$\frac{1}{(2\pi)^{(p+q)/2}|\Sigma|^{1/2}} = \frac{1}{(2\pi)^{p/2}|\Sigma/\Sigma_{22}|^{1/2}} \frac{1}{(2\pi)^{q/2}|\Sigma_{22}|^{1/2}} \quad (12)$$

Conditional distribution:

$$\mu_{1|2} = \mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(x_2 - \mu_2) \quad (13)$$

$$\Sigma_{1|2} = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21} \quad (14)$$

Marginalization:

$$\mu_2^m = \mu_2 \quad (15)$$

$$\Sigma_{22}^m = \Sigma_{22} \quad (16)$$

Canonical Parameters:

$$\Lambda = \Sigma^{-1} \quad (17)$$

$$\eta = \Sigma^{-1}\mu \quad (18)$$

Conditioning:

$$\eta_{1|2}^c = \eta_1 - \Lambda_{12}x_2 \quad (19)$$

$$\Lambda_{1|2}^c = \Lambda_{11} \quad (20)$$

Marginalization:

$$\eta_2^m = \eta_2 - \Lambda_{21}\Lambda_{11}^{-1}\eta_1 \quad (21)$$

$$\Lambda_2^m = \Lambda_{22} - \Lambda_{21}\Lambda_{11}^{-1}\Lambda_{12} \quad (22)$$

Factor Analysis:

$$x_n \sim N(0, I) \quad (23)$$

$$y_n \sim N(\mu + \Lambda x_n, \Psi) \quad (24)$$

Write

$$y_n = \mu + \Lambda x_n + \epsilon_n; \quad \epsilon_n \sim N(0, \Psi) \quad (25)$$

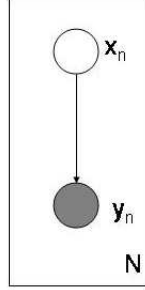
Maximum Likelihood (ML) estimation:

Let

$$\mathcal{D} = \{y_n\}_{n=1}^N; \quad \theta = (\mu, \Lambda, \Psi) \quad (26)$$

$$\ell(\theta|\mathcal{D}) = \sum_{n=1}^N \log p(y_n|\theta) \quad (27)$$

$$= -\frac{N}{2} \log |\Lambda\Lambda^T + \Psi| - \frac{1}{2} \left\{ \sum_{n=1}^N (y_n - \mu)^T (\Lambda\Lambda^T + \Psi)^{-1} (y_n - \mu) \right\} \quad (28)$$

Figure 1: Graphical model for multinomial density estimation with Dirichlet prior with hyper-parameter α .

$$E(y_n) = \mu + \Lambda E(x_n) + E(\epsilon_n) = \mu \quad (29)$$

$$\text{Var}(y_n) = E(\Lambda x_n + \epsilon_n)(\Lambda x_n + \epsilon_n)^T \quad (30)$$

$$= E(\Lambda x_n x_n^T \Lambda^T + \epsilon_n x_n^T \Lambda^T + \Lambda x_n \epsilon^T + \epsilon_n \epsilon_n^T) \quad (31)$$

$$= \Lambda \Lambda^T \quad (32)$$

Important formulas:

$$E(Y) = E(E(Y|X)) \quad (33)$$

$$\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X)) \quad (34)$$

So,

$$E(y_n|x_n) = \mu + \Lambda x_n \quad (35)$$

$$\text{Cov}(x_n, y_n) = E[x_n(\mu + \Lambda x_n + \epsilon_n - \mu)^T] = \Lambda^T \quad (36)$$

Thus, the overall mean and covariance matrix is given by

$$\mu = \begin{pmatrix} 0 \\ \mu \end{pmatrix} \quad (37)$$

$$\Sigma = \begin{pmatrix} I & \Lambda^T \\ \Lambda & \Lambda \Lambda^T + \Psi \end{pmatrix} \quad (38)$$

EM algorithm:

Write complete log likelihood (using the trace trick):

$$\ell(\theta|\mathcal{D}) = -\frac{N}{2} \log |\Psi| - \frac{1}{2} \sum_{n=1}^N x_n^T x_n - \frac{1}{2} \left\{ \sum_{n=1}^N (y_n - \Lambda x_n)^T \Psi^{-1} (y_n - \Lambda x_n) \right\} \quad (39)$$

$$= -\frac{N}{2} \log |\Psi| - \frac{1}{2} \sum_{n=1}^N \text{tr}(\mathbf{x}_n \mathbf{x}_n^T) - \frac{1}{2} \left\{ \sum_{n=1}^N \text{tr}[(\mathbf{y}_n - \Lambda \mathbf{x}_n)(\mathbf{y}_n - \Lambda \mathbf{x}_n)^T \Psi^{-1}] \right\} \quad (40)$$

Use the following formulas to obtain the EM algorithm:

$$\frac{\partial}{\partial A} \log |A| = A^{-T} \quad (41)$$

$$\frac{\partial}{\partial A} \text{tr}(BA) = B^T \quad (42)$$

$$\frac{\partial}{\partial A} \text{tr}(BA^T CA) = 2CAB \quad (43)$$