

**Stat 260/CS 294 Homework Assignment 2** (due March 1)

1. Find a transform of  $\theta$ ,  $\eta = g(\theta)$ , such that the Fisher information  $I(\eta)$  is constant (and therefore the Jeffreys prior is constant) for:
  - (a) the Poisson distribution,  $\text{Poi}(\theta)$ ;
  - (b) the gamma distribution,  $\text{Ga}(\alpha, \theta)$ , with  $\alpha = 1, 2, 3$ ;
  - (c) the binomial distribution,  $\text{Bin}(n, \theta)$ ; and
  - (d) the Maxwell distribution,  $\text{Max}(\theta)$ :  $p(x|\theta) \propto \theta^{3/2} x^2 e^{-\frac{\theta x}{2}}$ ,  $x \geq 0, \theta > 0$ .
  
2. Consider  $n$  random variables  $x_1, \dots, x_n$ , such that the first  $\xi$  of these variables has an  $\text{Exp}(\eta)$  distribution and the  $n - \xi$  other have an  $\text{Exp}(c\eta)$  distribution, where  $c$  is known and  $\xi$  takes its values in  $\{1, 2, \dots, n - 1\}$ .
  - (a) Let  $z = (z_2, \dots, z_n)$  where  $z_i = x_i/x_1$ . Using Bayes' theorem, give the shape of the posterior distribution  $p(\xi|x)$  when we use the prior  $p(\xi, \eta) = p(\xi)$  and show that it only depends on  $z$ .
  - (b) Directly (without using Bayes' theorem) show that the distribution of  $z$ ,  $p(z|\xi, \eta)$ , only depends on  $\xi$ ; i.e.,  $p(z|\xi, \eta) = p(z|\xi)$ .
  - (c) Show that the posterior distribution  $p(\xi|x)$  computed in (a) cannot be written as a product of the posterior distribution for  $p(z|\xi)$  as computed in (b) and  $p(\xi)$ , whatever  $p(\xi)$  is, even though we showed in part (a) that  $p(\xi|x)$  only depends on  $z$ . How do you explain this?
  - (d) Show that the paradox does not occur when we use the scale invariant prior  $p(\xi, \eta) = p(\xi)\eta^{-1}$ .

3. Consider  $n$  observations from a multinomial distribution where each observation belongs to one of 3 categories, with probabilities  $\theta_1, \theta_2$ , and  $1 - \theta_1 - \theta_2$ . The probability of observing  $r_1, r_2$  and  $n - r_1 - r_2$  observations from the three categories is given by:  $p(r_1, r_2|n, \theta_1, \theta_2) = \frac{n!}{r_1!r_2!(n-r_1-r_2)!} \theta_1^{r_1} \theta_2^{r_2} (1 - \theta_1 - \theta_2)^{n-r_1-r_2}$  for  $0 \leq r_1 + r_2 \leq n$ .

We are interested in the ratio  $\phi = \frac{\theta_1}{\theta_2}$  while we treat  $\lambda = \theta_2$  as a nuisance parameter. Assume that the joint distribution of  $(\lambda, \phi)$  is asymptotically normal.

- (a) Derive the covariance matrix of the asymptotic posterior distribution of  $(\lambda, \phi)$ . Also derive the asymptotic marginal standard deviation of  $\phi$  and the asymptotic conditional standard deviation of  $\lambda$  given  $\phi$ .
- (b) Show that the conditional reference prior is

$$\pi(\lambda|\phi) \propto (\lambda(1 + \phi))^{-\frac{1}{2}} (1 - \lambda(1 + \phi))^{-\frac{1}{2}}, \quad 0 < \lambda < \frac{1}{1 + \phi}$$

- (c) Show that the marginal reference prior is

$$\pi(\phi) \propto \phi^{-\frac{1}{2}} (1 + \phi)^{-1}$$

- (d) Derive the marginal posterior

$$\pi(\phi|r_1, r_2, n)$$

What does the form of the posterior say about the dependence of  $\phi$  on the observations?