

**CS294-179 Network Structure and Epidemics**  
**Fall 2020**  
**Homework #5**  
**Due Mo, Nov. 9**

## 1 Polya Urn

[5 Points]

Recall that a polya urn with  $R$  initial red, and  $B$  initial blue balls is a process where  $R$  and  $B$  are updated according to the following rule:

- Set  $R_0 = R$  and  $B_0 = B$
- At time  $t$ , draw a random ball from all  $B_t + R_t$  balls, and raise the  $R_t$  by one if the ball is red - otherwise raise  $B_t$  by one. Denote the new number of red balls by  $R_{t+1}$  and the new number of blue balls by  $B_{t+1}$ .

Set  $X_t = 1$  if the ball drawn at time  $t$  is red, and  $X_t = 0$  otherwise. Fix a set  $I_R \subset [n]$  and it's complement,  $I_B$ . We have seen that the probability of drawing a sequence  $X_1, \dots, X_n$  such that  $X_i = 1$  for  $i \in I_R$  and  $X_i = 0$  for  $i \in I_B$  does only depend on the sizes  $n_1$  and  $n_2$  of  $I_1$  and  $I_2$  and is equal to

$$\begin{aligned} \Pr(X_i = 1 \text{ for } i \in I_R \text{ and } X_i = 0 \text{ for } i \in I_B) \\ = \frac{R(R+1) \dots (R+n_1-1) \times B(B+1) \dots (B+n_2-1)}{(B+R)(B+R+1) \dots (B+R+n-1)}. \end{aligned}$$

In this exercise, you will prove the following theorem

**Theorem 1** *The probability of the sequence  $X_1, \dots, X_n$  draws from the Poly Urn with initially  $R$  red and  $B$  blue can equivalently be calculated by first drawing  $p \sim \beta(R, B)$  and then choosing  $X_i$  iid with distribution  $Be(p)$ .*

Here  $\beta(R, B)$  is the probability distribution on  $[0, 1]$  that has the probability density function  $\frac{1}{Z} x^{R-1} (1-x)^{B-1}$ , where  $Z = \int_0^1 x^{R-1} (1-x)^{B-1}$ .

1. Let  $A$  and  $B$  non-negative integers. Use integration by parts to calculate the integral

$$\int_0^1 x^A (1-x)^B dx.$$

2. Let  $R, B, n_1$  and  $n_2$  be non-negative integers, and assume that  $X \sim \beta(A, B)$ . Calculate first the normalization factor  $Z = \int_0^1 x^{R-1} (1-x)^{B-1}$  for the  $\beta$ -distribution, and then the expectation of  $X^{n_1} (1-X)^{n_2}$ .
3. Use the result from 2 to prove the theorem.

## 2 Degrees for preferential attachment

[7 Points]

Let  $D_n$  be the degree of a random vertex in the preferential attachment graph on  $n$  nodes (note that it has two sources of randomness - the randomness from choosing a vertex, and that of the preferential attachment graph). For the various version of preferential attachment discussed in the lectures (independent, conditional, and sequential), the random variable  $D_n$  converges in distribution to a random variable  $D$  whose distribution is given by

$$\Pr(D = k) = \frac{2m(m+1)}{k(k+1)(k+2)}.$$

A different way of obtaining the distribution of  $D$  proceeds by proving that Preferential attachment has a weak local limit, and then establishing the degree distribution for the root of the limiting rooted graph. In that approach, the variable  $D$  appears naturally as a sum of  $m$  and a mixed Poisson random variable  $q$ :

$$D = m + q \quad \text{where} \quad q \sim Poi\left(\gamma \frac{1-x}{x}\right)$$

where  $\gamma \sim \Gamma(m, 1)$  and  $x = \sqrt{u}$  with  $u$  being a uniform random variable in  $[0, 1]$ . In this exercise, you will prove that the two are indeed equivalent.

1. Let  $A$  be a non-negative integer. Use integration by parts to calculate the integral

$$\int_0^{\infty} x^A e^{-x} dx$$

2. Let  $A$  and  $k$  be non-negative integers, and let  $c > -1$ . A random variable  $X$  has distribution  $\Gamma(A, 1)$  if its probability density function is equal to  $\frac{1}{Z} x^{A-1} e^{-x}$ , where  $Z$  is a normalization constant. Calculate first  $Z$ , and then the expectation of  $X^k e^{-cX}$ .
3. Calculate the probability that the random variable  $q$  above takes the value  $k$ , conditioned on  $\gamma$  and  $x$ .
4. By first taking expectation with respect to the random variable  $\gamma$ , and then with respect to  $u = x^2$ , get the probability that  $q = k$ .
5. Relate this result to the explicit formula given above and show that the two give the same probability distribution.

### 3 Infection digraph

[6 Points]

Recall that the exponential distribution with rate  $\gamma$  is the probability distribution on  $\mathbb{R}_+$  with density  $\rho(T) = \gamma e^{-\gamma T}$ .

For a graph  $G$ , recovery rate  $\gamma$  and infection rate  $\tau$ , the infection digraph of an SIR model on  $G$  is determined as follows: for all vertices  $i$ , draw a recovery clock  $T_i \sim \text{exp}(\gamma)$  and for all edges  $\{i, j\}$  in  $E(G)$  draw two infection clocks  $T_{ij} \sim \text{exp}(\tau)$  and  $T_{ji} \sim \text{exp}(\tau)$ , all independently of each other. The infection digraph  $D_{SIR}$  is then defined by putting an oriented edge  $ij$  from  $i$  to  $j$  whenever  $\{i, j\}$  is an edge in  $G$  and  $T_{ij} \leq T_i$ . We set  $X_{ij} = 1$  if this is the case, and  $X_{ij} = 0$  otherwise. Note that the random variables  $X_{ij}$  are independent if we condition on the recovery times  $\{T_i\}_{i \in V(G)}$ .

1. Let  $p_T = 1 - e^{-\tau T}$ . Use elementary calculations involving the exponential distribution  $\text{exp}(\tau)$  to show that conditioned on the recovery times,  $X_{ij} \sim \text{Be}(p_{T_i})$  whenever  $\{i, j\}$  is an edge in  $G$ .
2. For a vertex  $i$ , let  $N(i)$  be the set of neighbors of  $i$  in  $G$ . Show that for a vertex  $i$  of degree  $d_i$  in  $G$ , the in-degree in  $D_{SIR}$ ,  $d_i^- = \sum_{j \in N(i)} X_{ji}$ , has distribution  $\text{Bin}(d_i, p)$ , where  $p = \frac{\tau}{\gamma + \tau}$ . Write down the generating function  $G_i^-(\eta)$  for the random variable  $d_i^-$ .
3. Show that conditioned on  $T_i$ , the out-degree of  $i$  in  $D_{SIR}$ ,  $d_i^+ = \sum_{j \in N(i)} X_{ij}$ , has distribution  $\text{Bin}(d_i, p_{T_i})$ . Write down the generating function  $G_i^+(\eta|T_i)$  for the random variable  $d_i^+$  conditioned on  $T_i$ .

Note that the generating function of the unconditioned random variable  $d_i^+$  is just the expectation of  $G_i^+(\eta|T_i)$  with respect to  $T_i$ ,

$$G_i^+(\eta) = \gamma \int G_i^+(\eta|T) e^{-\gamma T} dT.$$

(this is just an observation, you don't need to simplify this further for the moment).

4. Specializing to  $G = K_n$  and  $\tau = \beta/n$ , calculate  $G^-(\eta)$  and  $G^+(\eta|T_i)$  in the limit  $n \rightarrow \infty$  (if you don't make a mistake, you should get the generating functions of two Poisson random variables; also, since  $K_n$  is regular, the result will not depend on  $i$ , so we dropped the index  $i$ ). What are the rates  $c^-$  and  $c^+(T_i)$  of these two Poisson random variables?
5. Using the above, it is easy to calculate the generating function  $G^+$  of the unconditioned out-degree  $d_i^+$ , since by dominated convergence it converges to the expectation of the limit you just calculated. Calculate this expectation, and express the limiting generating function of  $d_i^+$  as an explicit function of  $R_0 = \beta/\gamma$  and  $\eta$ . Compare the expression you get to that of a geometric random variable.
6. (Bonus) Explicitly solve the equation  $\theta + G^+(1 - \theta) = 1$  for the survival probability  $\theta$  of a birth process with off spring distribution  $d^+$  (in the limit  $n \rightarrow \infty$ , where, as discussed in class, it is the survival probability of the infection). Compare the results to those obtain from  $G^-$ , both for  $R_0 = 1 + \epsilon$  for small  $\epsilon$  (in which case I'd like you to compare the rate at which  $\theta \rightarrow 0$  as  $\epsilon \rightarrow 0$ ) as well as for large  $R_0$  (in which case I'd like you to compare the rate at which  $\eta = 1 - \theta \rightarrow 0$  as  $R_0 \rightarrow \infty$ ).